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SOLUTIONS TO ASSIGNMENT - 8

1. Differential Entropy

(a) Exponential distribution

$$h(f) = - \int_0^{\infty} \lambda e^{-\lambda x} (\ln \lambda - \lambda x) dx = -\ln \lambda + 1 \text{ nats}$$

$$= \log \frac{e}{\lambda} \text{ bits}$$

(b) Laplace density

$$h(f) = - \int_{-\infty}^{\infty} \frac{1}{2} \lambda e^{-\lambda |x|} \left[ \ln \frac{1}{2} + \ln \lambda - \lambda |x| \right] dx$$

(even fn)

$$= -\ln \frac{1}{2} - \ln \lambda + 1$$

$$= \ln \frac{2e}{\lambda} \text{ nats} = \log \frac{2e}{\lambda} \text{ bits}$$

(c) Sum of two normal distributions.

$$\text{Now } X_1 + X_2 \sim \mathcal{N}(\mu_1 + \mu_2, \sigma_1^2 + \sigma_2^2)$$

$$h(f) = \frac{1}{2} \log 2\pi e (\sigma_1^2 + \sigma_2^2) \text{ bits}$$

2)  $X_1 \sim \mathcal{N}(0, K_1)$   $X_2 \sim \mathcal{N}(0, K_2)$

$$\Pr(\theta=1) = \lambda \quad \Pr(\theta=2) = \bar{\lambda}$$

$$\text{Now let } \boxed{Z = X_{\theta}} \Rightarrow K_Z = \lambda K_1 + \bar{\lambda} K_2$$

Now  $h(Z) \leq \frac{1}{2} \log (2\pi e)^n |\lambda K_1 + \bar{\lambda} K_2|$

$\rightarrow$   $Z$  is not multivariate normal. (Gaussian maximises entropy)

But  $h(Z) \geq h(Z|e) = \lambda h(X_1) + \bar{\lambda} h(X_2)$

$$= \frac{\lambda}{2} \log (2\pi e)^n |K_1| + \frac{\bar{\lambda}}{2} \log (2\pi e)^n |K_2|$$

$$= \frac{1}{2} \log \left\{ (2\pi e)^n |K_1|^{\lambda} (2\pi e)^n |K_2|^{\bar{\lambda}} \right\}$$

$$= \frac{1}{2} \log \left\{ (2\pi e)^n |K_1|^{\lambda} |K_2|^{\bar{\lambda}} \right\}$$

Thus  $\frac{1}{2} \log (2\pi e)^n |\lambda K_1 + \bar{\lambda} K_2| \geq h(Z) \geq \frac{1}{2} \log \left\{ (2\pi e)^n |K_1|^{\lambda} |K_2|^{\bar{\lambda}} \right\}$

$$\Rightarrow |\lambda K_1 + \bar{\lambda} K_2| \geq |K_1|^{\lambda} |K_2|^{\bar{\lambda}}$$

3)  $\begin{pmatrix} x \\ y \end{pmatrix} \sim N\left(0, \begin{bmatrix} \sigma^2 & \rho\sigma^2 \\ \rho\sigma^2 & \sigma^2 \end{bmatrix}\right)$

$$h(x, y) = \frac{1}{2} \log (2\pi e)^2 |K| = \frac{1}{2} \log \left\{ (2\pi e)^2 \sigma^4 (1-\rho^2) \right\}$$

~~$x, y$~~   $x \sim N(0, \sigma^2) \quad y \sim N(0, \sigma^2)$

$$\Rightarrow h(x) = h(y) = \frac{1}{2} \log 2\pi e \sigma^2$$

$$I(x; y) = h(x) + h(y) - h(x, y) = -\frac{1}{2} \log (1-\rho^2)$$

$$(a) \rho = 1 \Rightarrow I(X; Y) = \infty$$

Since  $\rho = 1$   $X = Y$ . Hence mutual information  $\rightarrow \infty$

$$(b) \rho = 0. \quad I(X; Y) = 0 \text{ since } X \text{ \& } Y \text{ are independent}$$

and hence  $h(Y|X) = h(Y)$  so  $I(X; Y) = 0$

$$(c) \rho = -1 \Rightarrow X = -Y. \text{ Again from (a) } I(X; Y) = \infty$$

(sign doesn't matter)

$$4) \quad Y = X + Z$$

$\Rightarrow$  distribution of  $Y$  = convolution of densities of  $X + Z$  (since they are independent)

$$P_Y(y) = \begin{cases} \frac{1}{a} \left( y + \left( \frac{1+a}{2} \right) \right) & -\left( \frac{1+a}{2} \right) \leq y \leq -\left( \frac{1-a}{2} \right) \\ 1 & -\left( \frac{1-a}{2} \right) \leq y \leq \left( \frac{1-a}{2} \right) \\ \frac{1}{a} \left( -y + \left( \frac{1+a}{2} \right) \right) & \left( \frac{1-a}{2} \right) \leq y \leq \left( \frac{1+a}{2} \right) \end{cases}$$

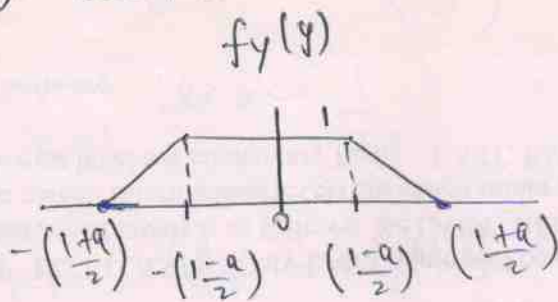
and for  $a \geq 1$

$$P_Y(y) = \begin{cases} \frac{1}{a} \left( y + \left( \frac{1+a}{2} \right) \right) & -\left( \frac{a+1}{2} \right) \leq y \leq -\left( \frac{a-1}{2} \right) \\ \frac{1}{a} & -\left( \frac{a-1}{2} \right) \leq y \leq \left( \frac{a-1}{2} \right) \\ \frac{1}{a} \left( -y + \left( \frac{1+a}{2} \right) \right) & \left( \frac{a-1}{2} \right) \leq y \leq \left( \frac{a+1}{2} \right) \end{cases}$$

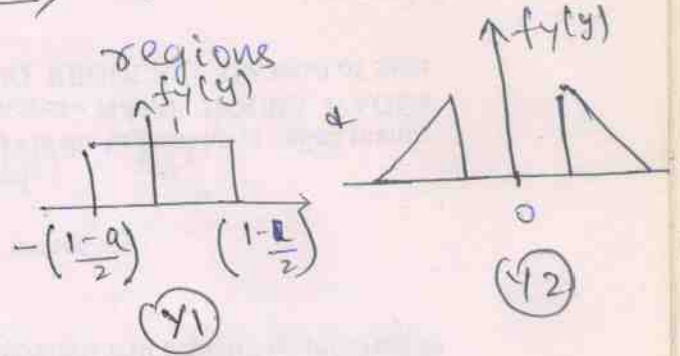
$a = 1 \Rightarrow Y$  is triangular over  $(-1, 1)$



Case 1)  $a \leq 1$



$\Rightarrow$  divide into two



Now let

$$\lambda = \begin{cases} 1 & \text{if } Y = Y_1 \\ 2 & \text{if } Y = Y_2 \end{cases}$$

Thus  $X = f(Y)$

$$\Rightarrow h(X|Y) = 0$$

$$\begin{aligned} \text{Now } h(\lambda, Y) &= h(Y) + h(\lambda|Y) = h(Y) \\ &= h(\lambda) + h(Y|\lambda) \end{aligned}$$

$$\text{Thus } h(Y) = h(\lambda) + h(Y|\lambda)$$

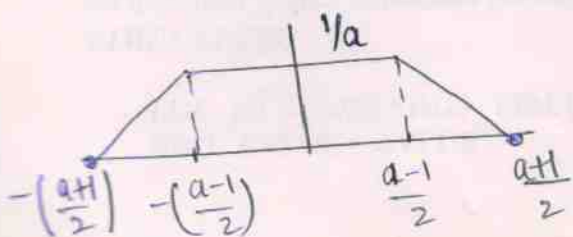
$$= h(1-\alpha) + (1-\alpha)h(Y_1) + \alpha h(Y_2)$$

$$= h(1-\alpha) + (1-\alpha) \ln \frac{1}{1-\alpha} + \alpha \left( \frac{1}{2} + \ln a \right)$$

$$= \frac{\alpha}{2} \text{ nats}$$

Triangular distribution

Case 2)  $a \geq 1$



$$\text{Now } P(\lambda=1) = \frac{a-1}{a} \left\{ = P(Y_1) \right\}$$

$$\begin{aligned} h(Y) &= h(\lambda) + h(Y|\lambda) \\ &= h\left(\frac{1}{a}\right) + \left(\frac{a-1}{a}\right) \ln(a-1) \\ &\quad + \frac{1}{a} \left( \frac{1}{2} + \ln 1 \right) \end{aligned}$$

$$= \frac{1}{a} \ln a + \left(\frac{a-1}{a}\right) \ln \frac{a}{a-1} + \frac{a-1}{a} \ln(a-1) + \frac{1}{2a}$$

$$= \ln a + \frac{1}{2a}$$

$$I(x; Y) = h(Y) - h(Y|X) = h(Y) - h(Z)$$

$$h(Z) = \ln a$$

$$\text{Thus } I(x; Y) = \begin{cases} \frac{a}{2} - \ln a & a \leq 1 \\ \frac{1}{2a} & a \geq 1 \end{cases}$$

(b)  $a=1 \Rightarrow$  range of  $Y$  is  $(-1, 1)$

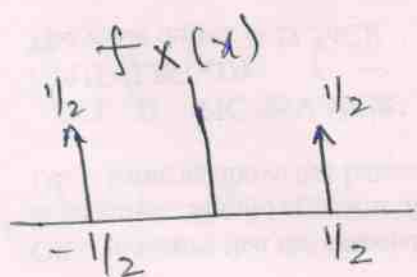
$$I(x; Y) = h(Y) - h(Y|X) = h(Y) - h(Z)$$

$$= h(Y) - \ln 1 = h(Y)$$

$h(Y)$  is maximum when it is uniform over  $(-1, 1)$

$$\Rightarrow I(x; Y) = h(Y) = \ln 2$$

what  $x$  makes  $Y$  uniform? The convolution of



$f_X(x) * f_Z(z)$  is uniform.

5) Two look Gaussian:

The input distribution that maximises capacity is  $X \sim N(0, P)$ .

$$\begin{aligned} C &= \max I(X; Y_1, Y_2) \\ &= h(Y_1, Y_2) - h(Y_1, Y_2 | X) = h(Y_1, Y_2) - h(Z_1, Z_2 | X) \\ &= h(Y_1, Y_2) - h(Z_1, Z_2) \quad (Z_1, Z_2 \text{ independent of } X) \end{aligned}$$

$$h(Z_1, Z_2) = \frac{1}{2} \log(2\pi e)^2 |K_Z| = \frac{1}{2} \log(2\pi e)^2 N^2(1-p^2)$$

Now

$$(Y_1, Y_2) \sim N\left(0, \begin{bmatrix} P+PN & P+PN \\ P+PN & P+PN \end{bmatrix}\right)$$

$$h(Y_1, Y_2) = \frac{1}{2} \log(2\pi e)^2 (N^2(1-p^2) + 2PN(1-p))$$

$$\text{Thus } C = \frac{1}{2} \log\left(1 + \frac{2P}{N(1+p)}\right)$$

$$(a) \quad p=1 \quad C = \frac{1}{2} \log\left(1 + \frac{P}{N}\right) \quad \text{since } Y_1 = Y_2$$

$$(b) \quad p=0 \quad C = \frac{1}{2} \log\left(1 + \frac{2P}{N}\right)$$

$$(c) \quad p=-1 \quad C = \infty. \quad \text{Since } Z_1 = -Z_2 \text{ and adding } (Y_1 + Y_2) \text{ we can cancel out noise.}$$



6) Multipath Gaussian:

$$Y = (a_1 + a_2)X + Z_1 + Z_2$$

$$K_Z = \begin{bmatrix} \sigma^2 & p\sigma^2 \\ p\sigma^2 & \sigma^2 \end{bmatrix} \Rightarrow Z_1 + Z_2 \sim N(0, \sigma_Z^2)$$

$$\begin{aligned} \sigma_Z^2 &= (\sigma^2 + p\sigma^2) + (\sigma^2 + p\sigma^2) \\ &= 2\sigma^2(1+p) \end{aligned}$$

Thus

$$C = \max_{I(X; Y)} = \frac{1}{2} \log_2 \left( 1 + \frac{(a_1 + a_2)^2 P}{2\sigma^2(1+p)} \right)$$

a)  $p=0 \Rightarrow C = \frac{1}{2} \log_2 \left( 1 + \frac{(a_1 + a_2)^2 P}{2\sigma^2} \right)$

b)  $p=1 \Rightarrow C = \frac{1}{2} \log_2 \left( 1 + \frac{(a_1 + a_2)^2 P}{4\sigma^2} \right)$  since  $Z_1 = Z_2$

c)  $p=-1 \Rightarrow C = \frac{1}{2} \log_2(1 + \infty) = \infty$  since  $Z_1 = -Z_2$  (cancel)

7) Parallel Gaussian channels:

Given  $\beta_1 P_1 + \beta_2 P_2 \leq \beta$

(a) Note  $C_i = \frac{1}{2} \log \left( 1 + \frac{P_i}{N_i} \right) = \frac{1}{2} \log \left( 1 + \frac{\beta_i P_i}{\beta_i N_i} \right)$

New Noise powers:  $\beta_1 N_1$  and  $\beta_2 N_2$ .

Ideal solution is waterfilling.

In waterfilling solution parallel channels stop acting like single channel when  $|P| \geq |N_1 - N_2|$ . Applying this here.

$$\beta \geq |\beta_1 N_1 - \beta_2 N_2|$$

(b) Now  $\beta_1 N_1 = 3$   $\beta_2 N_2 = 4$

using water filling  $\beta_1 P_1 = 5.5$   $\beta_2 P_2 = 4.5$

thus optimum power  $P_1 = 5.5$   $P_2 = 2.25$

and  $C = C_1 + C_2 = \frac{1}{2} \log\left(1 + \frac{5.5}{3}\right) + \frac{1}{2} \log\left(1 + \frac{2.25}{2}\right)$   
 $= 1.29$  bits.