

EE 6110 Adaptive Signal Processing

Prerequisites: Probability Theory (Necessary)
Estimation Theory (Optional)

Textbook Fundamentals of Adaptive filtering
by Ali Sayed Wiley Indian Edition

Syllabus

- Optimal Estimation
- Linear Estimation
- Steepest Descent Algorithms
- Stochastic Gradient Algorithms \rightarrow adaptive filters.
- Performance of adaptive filters.

Evaluation

Assignments	10%
Quiz 1 & 2	20+20
Final Exam	50%

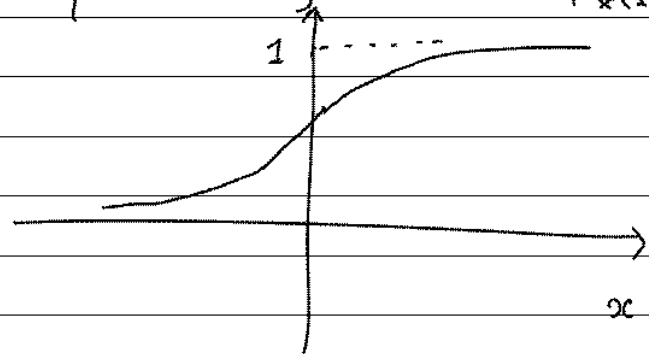
Review of Probability Basics

X is a Random Variable.

Cumulative distribution function of X (cdf)

$$F_x(x) = \Pr \{ X \leq x \}$$

\downarrow
increasing function of x .



$$F_x(-\infty) = 0$$

$$F_x(\infty) = 1$$

Note $\Pr \{ a < X \leq b \} = F_x(b) - F_x(a)$

Probability density function (pdf) of X

$$f_x(x) = \frac{d}{dx} F_x(x) \quad (\text{if derivative exists})$$

$$= \lim_{\Delta x \rightarrow 0} \frac{F_x(x + \Delta x) - F_x(x)}{\Delta x}$$

$$\Pr \{ a < X \leq b \} = \int_a^b f_x(x) dx$$

$$f_x(x) \geq 0 \quad \text{for all } x.$$

Conditional distribution

Let X be random variable

M be an event (eg. $M = \{a < X \leq b\}$)

Conditional distribution of X given M

$$F_x(x|M) = \Pr\{X \leq x | M\}$$

↘ conditional probability

$$= \frac{\Pr\{X \leq x, M\}}{\Pr\{M\}}$$

Conditional density of X given M

$$f_x(x|M) = \frac{d}{dx} F_x(x|M)$$

Example:

X be rv denoting life time of a machine.

Let $M = \{X > t_0\}$

We know $F_x(x)$, $f_x(x)$

Want to find $F_x(x|M)$, $f_x(x|M)$

$$F_X(x|M) = \Pr\{X \leq x | M\}$$

$$= \frac{\Pr\{X \leq x, X > t_0\}}{\Pr\{X > t_0\}}$$

$$F_X(x|M) = \begin{cases} 0 & ; x \leq t_0 \\ \frac{F_X(x) - F_X(t_0)}{1 - F_X(t_0)} & ; x > t_0 \end{cases}$$

Conditional pdf

$$f_X(x|M) = \begin{cases} 0 & ; x \leq t_0 \\ \frac{f_X(x)}{1 - F_X(t_0)} & ; x > t_0 \end{cases}$$

x $\xrightarrow{\hspace{10em}}$ x

Expected Values

X is random variable
with pdf $f_X(x)$

$$\text{Mean of } X = E(X) = \int_{-\infty}^{\infty} x f_X(x) dx.$$

$$\text{Mean squared} = E(X^2) = \int_{-\infty}^{\infty} x^2 f_X(x) dx \quad \rightarrow \text{if it exists}$$

$$\geq 0$$

$$\text{Variance of } x = E(x^2) - \{E(x)\}^2 \geq 0$$

$$= E\left[(x - E(x))^2\right]$$

Suppose $g(\cdot)$ is any function.

$$E(g(x)) = \int_{-\infty}^{\infty} g(x) f_x(x) dx$$

x $\xrightarrow{\hspace{2cm}}$ x

Conditional Expectation.

M is an event

$$E(x|M) = \int_{-\infty}^{\infty} x f_x(x|M) dx$$

$$E(x^2|M) = \int_{-\infty}^{\infty} x^2 f_x(x|M) dx$$

$$\text{Conditional Variance} = E(x^2|M) - (E(x|M))^2$$

Example: let x be Gaussian with
pdf $\frac{1}{\sqrt{2\pi}} e^{-x^2/2}$

$$E(X) = 0$$

$$\text{Var}(X) = 1$$

$$M = \{X > 0\}$$

Need to compute

$$E(X|M) = ?$$

$$f_X(x|M)$$

$$\text{Var}(X|M) = ?$$

$$f_X(x|M) = \begin{cases} 2 f_X(x) & ; x > 0 \\ 0 & ; x < 0 \end{cases}$$

$$E(X|M) = \int_0^{\infty} 2x f_X(x) dx = \sqrt{2/\pi}$$

$$E(X^2|M) = \int_0^{\infty} 2x^2 f_X(x) dx$$

$$= \int_{-\infty}^{\infty} x^2 f_X(x) dx = 1$$

$$\text{Var}(X|M) = 1 - 2/\pi$$

x \longrightarrow

1/8
Joint Random Variables.

X and Y are two random variables

Joint distribution function $F_{X,Y}$

$$F_{XY}(x, y) = \Pr \{ X \leq x, Y \leq y \}$$

Joint density function f_{XY}

$$f_{XY}(x, y) = \frac{\partial^2}{\partial x \partial y} F_{XY}(x, y) \quad (\text{if the derivative exists})$$

We also have.

$$F_{XY}(x, y) = \int_{\alpha=-\infty}^x \int_{\beta=-\infty}^y f_{XY}(\alpha, \beta) d\alpha d\beta$$

Marginal distribution function $\rightarrow Y$ can take any value from $-\infty$ to ∞

$$F_X(x) = \Pr \{ X \leq x \}$$

$$= \Pr \{ X \leq x, Y \leq \infty \}$$

$$= F_{XY}(x, \infty)$$

Similarly $F_Y(y) = F_{XY}(\infty, y)$

Marginal

pdf of X $f_X(x) = \frac{d}{dx} F_X(x)$

$$= \frac{d}{dx} F_{XY}(x, \infty)$$

$$= \frac{d}{dx} \int_{\alpha=-\infty}^x \int_{\beta=-\infty}^{\infty} f_{XY}(\alpha, \beta) d\alpha d\beta$$

$$f_X(x) = \int_{-\infty}^{\infty} f_{XY}(x, \beta) d\beta$$

$$f_Y(y) = \int_{-\infty}^{\infty} f_{XY}(\alpha, y) d\alpha$$

x ————— x

Two random variables X and Y are independent if

$$F_{XY}(x, y) = F_X(x) F_Y(y)$$

for all x & y .

equivalently

$$f_{XY}(x, y) = f_X(x) f_Y(y)$$

x ————— x

Two random variables X and Y are called uncorrelated if

$$E(XY) = E(X) E(Y)$$

clearly from definitions

independent \Rightarrow uncorrelated.

uncorrelated $\not\Rightarrow$ independence. (except for jointly Gaussian random variables)

Property: If X and Y are uncorrelated then

$$\rightarrow \text{Var}(X+Y) = \text{Var}(X) + \text{Var}(Y)$$

Variance

Proof:
$$\text{Var}(X+Y) = E((X+Y)^2)$$

$$- \{E(X+Y)\}^2$$

$$\begin{aligned} E(X+Y) &= E(X^2 + Y^2 + 2XY) \\ &= E(X) + E(Y) \end{aligned}$$

$$- (E(X) + E(Y))^2$$

$$= E(X^2) + E(Y^2) + 2E(XY)$$

$$- (E(X))^2 - (E(Y))^2 - 2E(X)E(Y)$$

$$\text{Var}(X+Y) = \text{Var}(X) + \text{Var}(Y)$$

* ————— *

Two random variables are called orthogonal
(X & Y)

$$\text{if } E(XY) = 0$$

Note: If one of RV has zero mean
then orthogonality imply uncorrelated

* ————— *

Example.

Let Θ be RV
uniformly distributed in $[-\pi$ to $\pi]$

$$f_{\Theta}(\theta) = \begin{cases} \frac{1}{2\pi} & \text{if } \theta \in [-\pi, \pi] \\ 0 & \text{else} \end{cases}$$

$$\text{let } X = \cos \Theta$$

$$X^2 + Y^2 = 1$$

$$Y = \sin \Theta$$

X and Y are not
independent.

Are X and Y uncorrelated?

$$E(X) = E(\cos \Theta) = \int_{-\pi}^{\pi} \cos \theta \cdot f_{\Theta}(\theta) d\theta$$

$$= \int_{-\pi}^{\pi} \cos \theta \frac{1}{2\pi} d\theta = 0$$

$$E(Y) = 0$$

$$E(XY) = \int_{-\pi}^{\pi} \cos \theta \cdot \sin \theta \frac{1}{2\pi} d\theta$$

$$= \int_{-\pi}^{\pi} \frac{\sin 2\theta}{4\pi} d\theta = 0$$

$$E(XY) = 0 = E(X)E(Y)$$

X and Y are uncorrelated (also orthogonal)

Conditional distribution / density

X & Y are two RV.

$F_{XY}(x,y)$

joint cdf

$f_{XY}(x,y) \rightarrow$ joint pdf.

Conditional distribution of X given $Y \leq y$

$$F_X(x | Y \leq y) = \Pr\{X \leq x | Y \leq y\}$$

$$= \frac{\Pr\{X \leq x, Y \leq y\}}{\Pr\{Y \leq y\}}$$

$$F_X(x | Y \leq y) = \frac{F_{XY}(x,y)}{F_Y(y)}$$

x ————— x

Conditional density function of X given $Y = y$

$$f_X(x | y) = \frac{f_{XY}(x,y)}{f_Y(y)} \quad (\text{if it exists})$$

Interpretation

lim

$$\frac{\Pr\{x < X < x + \Delta x | y < Y \leq y + \Delta y\}}{\Delta x}$$

$\Delta x \rightarrow 0$

$\Delta y \rightarrow 0$

Δx

$$= \lim_{\substack{\Delta x \rightarrow 0 \\ \Delta y \rightarrow 0}} \frac{\Pr \{ x < X \leq x + \Delta x, y < Y \leq y + \Delta y \}}{\Delta x \Delta y - \underbrace{\Pr \{ y < Y \leq y + \Delta y \}}_{\Delta y}}$$

$$= \frac{f_{XY}(x, y)}{f_Y(y)} \rightarrow \text{conditional density of } X \text{ given } Y = y$$

$$= f_X(x|y)$$

~ ~ ~ ~ ~ x

Conditional Expectation of X given $Y = y$.

$$E(X | Y = y) = \int_{-\infty}^{\infty} x f_X(x|y) dx$$

↓

will be a function of y .

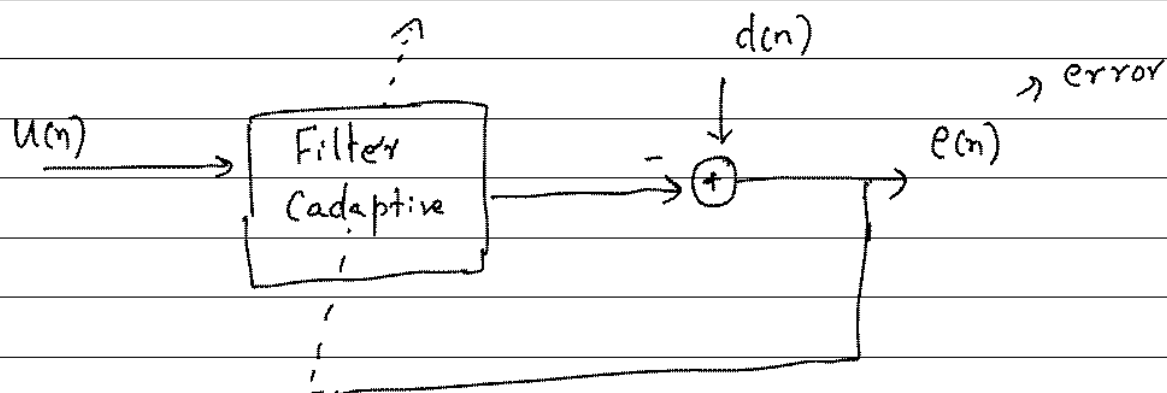
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Preview of Adaptive Filters.

Typical Problem:

$n \rightarrow$ time index

input signal $u(n)$ } sequence of
desired signal $d(n)$ } random variables



Goal: Minimize $E|e(n)|^2$

Adapt the filter based on $e(n)$

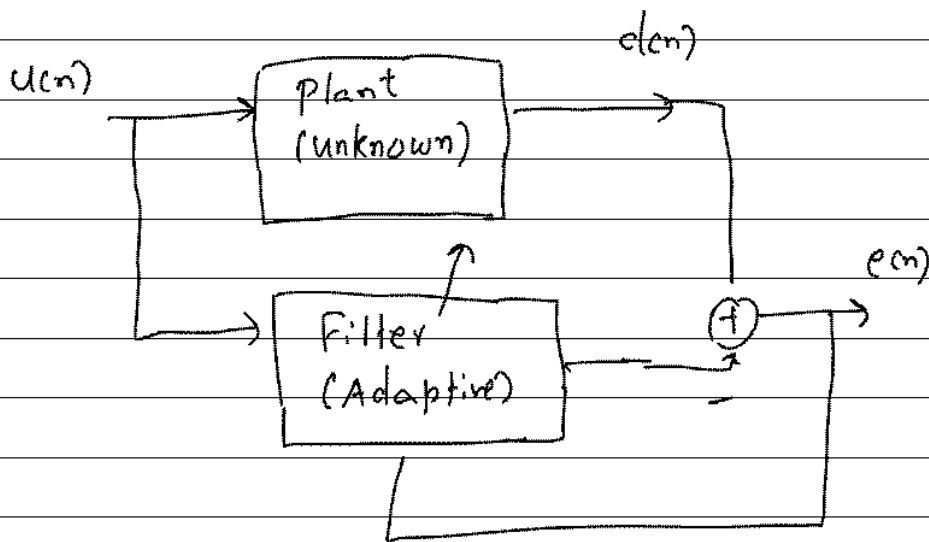
- If joint statistics of $u(n)$ and $d(n)$ are known then optimal filter can be computed in a single step

- Suppose joint statistics of $u(n)$ and $d(n)$ are not known or
Suppose statistics change with time } \Rightarrow Need for adaptive filtering

Questions: how do we learn statistics over time and will the adaptive filter converge to optimal filter?

Adaptive Filtering Applications

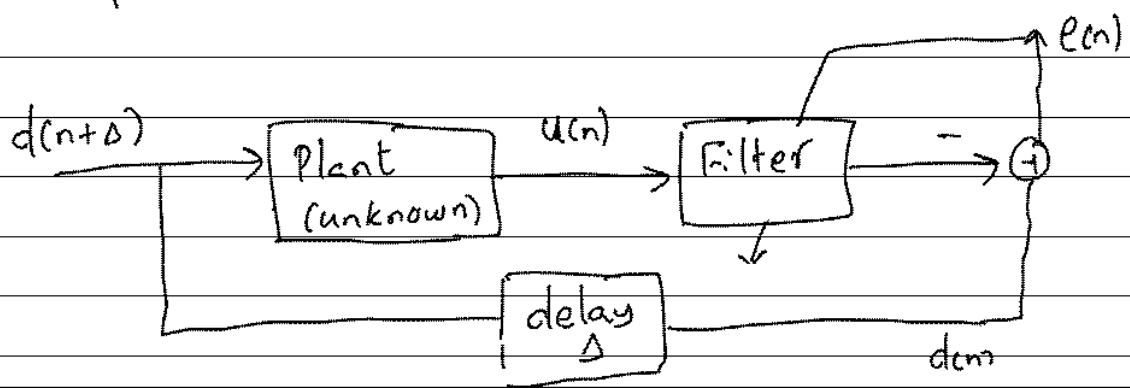
(a) System Identification



When $|e(n)|^2$ is small, adaptive filter response is close to plant response.

e.g.: channel estimation problem in communication

(b) Equalization problem



Here in the steady state when $e(n)^2$ is small,
filter response is close to inverse of
plant response

(c) other applications include
echo cancellation, interference cancellation

x ————— x